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## REVIEW

# Surveillance for antimicrobial resistant organisms: potential sources and magnitude of bias

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## SUMMARY

Surveillance has been recognized as a fundamental component in the control of antimicrobial-resistant infections. Although surveillance data have been widely published and utilized by researchers and decision makers, little attention has been paid to assessment of their validity. We conducted this review in order to identify and explore potential types and magnitude of bias that may influence the validity or interpretation of surveillance data. Six main potential areas were assessed. These included bias related to use of inadequate or inappropriate (1) denominator data, (2) case definitions, and (3) case ascertainment; (4) sampling bias; (5) failure to deal with multiple occurrences, and (6) those related to laboratory practice and procedures. The magnitude of these biases varied considerably for the above areas within different study populations. There are a number of potential biases that should be considered in the methodological design and interpretation of antimicrobial-resistant organism surveillance.

**Key words:** Emerging infections, epidemiology, microbiology, resistance to drugs.

## INTRODUCTION

The emergence of resistance to antimicrobial agents is a paramount contemporary healthcare issue [1]. Concern about emerging antimicrobial-resistant organisms (ARO) is increasing both in health professionals and the general public. Efforts to control resistance include, but are not limited to, infection prevention and control practices, prudent use of antimicrobials, and surveillance [2–7]. The Centers for Disease Control and Prevention (CDC) define

surveillance as ‘the ongoing systematic collection, analysis, and interpretation of health data essential to the planning, implementation, and evaluation of public health practice, closely integrated with the timely dissemination of these data to those who need to know’ [8]. There exist a large number of surveillance systems for detection of ARO worldwide that provide information on the occurrence of resistant infections in time and location, identify risk factors for their acquisition, and define their outcome [9–13]. Tracking these organisms and their determinants may aid policy-makers in their decisions regarding health-services funding allocation and guide research efforts into means of prevention and control. In addition, knowledge of species distribution and resistant

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profiles of infectious diseases is frequently touted as a means to improve appropriate antimicrobial utilization.

While it is evident that surveillance data may benefit understanding of the determinants and spread of antimicrobial-resistant infections and may aid clinicians with direct patient care decisions, the value of such information is predicated on its reliability and validity. Invalid ARO surveillance data risks wasting healthcare resources through misguided efforts and may result in patient harm through inappropriate use of antimicrobial agents. Despite the large number of surveillance systems worldwide and vast amounts of published data from resistance surveys, methodological issues surrounding ARO surveillance systems has rarely been the topic of investigation [14–21].

The objective of this review is to identify and explore potential types and magnitude of bias that may influence the validity or interpretation of surveillance data. The published English literature was reviewed in order to identify potential biases and to assess their magnitude in surveillance systems for ARO. Bias was defined as ‘systematic errors that may occur in collecting or interpreting data’ [22]. Information sources included Medline searches through the Ovid and PubMed platforms and Google Scholar searches between 1988 and 2008, examining bibliographies of selected articles, reviews, and the authors’ personal files and research databases. The review focused on surveillance for AROs. However, where data were lacking, examples from infectious diseases in general was used to support the underlying bias concepts.

Several potential biases or issues in interpretation were considered that may affect ARO surveillance systems and we categorized these into six groups as shown in Table 1. Each of these is further reviewed in the following sections.

## POTENTIAL BIASES

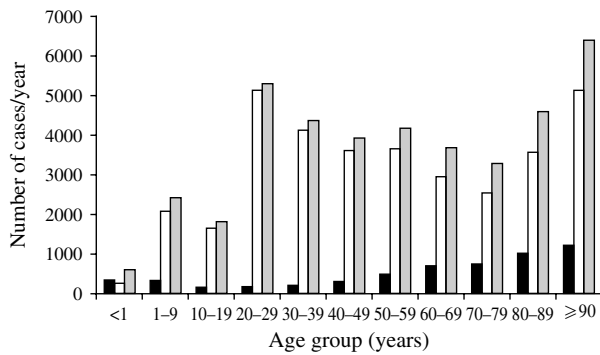
### Denominator data

How denominator data are utilized in surveillance data may have a significant influence on their interpretation, particularly surrounding the occurrence of an ARO infection. Hospital-acquired ARO infections may be used as an example to illustrate this point. By most definitions, these are infections that are not incubating or present at the time of admission, and are first identified more than 2 days after admission to hospital [23]. While an infection may be acquired

Table 1. *Potential biases in surveillance systems for antimicrobial resistance*

Bias	Explanation
Denominator data	Potential false attribution of risk related to incomplete data or interpretation of results if wrong denominator used
Case definition	If the method of case identification is not matched to surveillance objectives and/or reporting bias may result
Case ascertainment	Occurs with incomplete identification of all episodes that meet a case definition within the surveillance population or a failure to exclude those that do not meet that definition
Sampling bias	Occurs when sampling differs in some systematic way from the larger population of interest
Multiple counting	Bias that arises from counting a case more than once for the same episode of disease
Laboratory practice and procedures	Policies that direct laboratory testing protocols can introduce bias into surveillance systems. These include non-standardized testing, selective testing, rule-based reporting, and inadequate species level identification

within the first 2 days of hospital admission, it probably will not be detected by hospital surveillance in patients who have durations of admission <2 days. Should patients who are admitted for <2 days to hospital be included in the denominator for establishing nosocomial infection rates? If all admissions are used as the denominator, and a large number of patients are admitted for <2 days, then a potentially lower rate will be reported than if the denominator is restricted to those admitted for at least 2 days. In order to assess the degree of magnitude potentially associated with this we calculated the incidence of intensive-care unit (ICU)-acquired ARO bloodstream infections based on one of our previous works [24]. Using a denominator that restricts the population at risk to ICU admission duration of  $\geq 48$  h, the cumulative incidence and incidence density of ICU-acquired ARO bloodstream infection was 4.6/1000 admissions and 0.62/1000 patient-days, respectively. If all admissions are included, then these rates are substantially lower at 2.2/1000 admissions and 0.54/1000 patient-days, respectively.

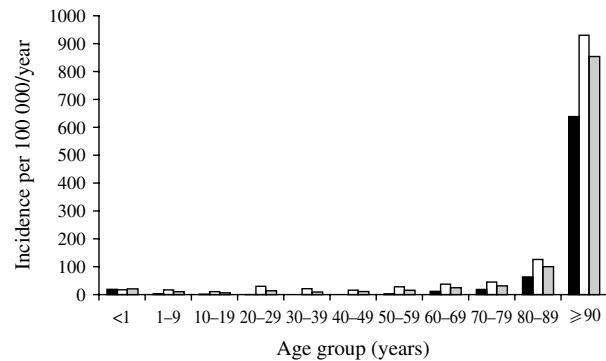


**Fig. 1.** Number of urinary tract infections in Calgary Health Region 2004/2005. Data are shown as number of cases. ■, Male; □, female; ▒, total cases. (Figure adapted from Laupland *et al.* [25].)

A second consideration with denominator data is that if it is not utilized at all, risk may be falsely attributed. An important example of this potential bias is exemplified by the occurrence of bacteriuria in a population [25]. Figure 1 shows the distribution of positive urinary-tract cultures in residents of the Calgary Health Region, Canada during 2004/2005. Based on these data as presented, one would conclude that women are at much higher risk than men, that there is a gradual increase in occurrence in men with advancing age, and that women have a decreasing occurrence through adulthood with a return to the same level in elderly women (Fig. 1). However, once these same data are expressed by inclusion of denominator data to express as population at risk, the conclusions are dramatically different (Fig. 2). Clearly the very old are at much higher risk, and the gender-related risk in this age group is no longer marked.

A third consideration is when the rate of resistant organisms is reported as a proportion of all isolates isolated. A false assessment of changes in the occurrence of resistance over time may occur when the denominator is changing due to some other factor. For example, if an antimicrobial is introduced into a setting with a particular pathogen, the result will be a decrease in the proportion of the isolation of susceptible bacteria to that antibiotic [26]. In such an instance even if the absolute number of resistant organism remains stable, the proportion (or occurrence) of resistance will be falsely assumed to be increasing [26]. Therefore, simply measuring the proportions of resistant bacteria to susceptible bacteria may result in misleading data if it is not interpreted in context.

While it is debatable as to which denominator may be preferred or 'correct' in these situations, the



**Fig. 2.** Incidence of urinary tract infections in Calgary Health Region 2004/2005. Data are shown as incidence per population. ■, Male; □, female; ▒, total cases. (Figure adapted from Laupland *et al.* [25].)

important point of consideration is that there is potential for misattribution of risk and occurrence depending on how denominator data are utilized.

#### Case definition

Bias may arise when a study case definition is not matched to surveillance objectives and/or reporting. Case definitions for infection may be syndromic (a constellation of clinical features) or may be definable based on a specific positive laboratory test such as a culture. This potential bias may be exemplified by urinary tract infections (UTIs). A diagnosis of a UTI involves integration of clinical symptoms (i.e. dysuria, frequency, fever, and/or pain), urinalysis results (pyuria, proteinuria, nitrates, and/or haematuria), and urine culture results [27]. While most UTIs are associated with a positive urine culture, this is not mandatory and a diagnosis may be established using a summation of other criteria. The use of an appropriate case definition is an important consideration in establishing and interpreting rates of resistance to antimicrobials in UTIs. If the surveillance objective is to identify rates of resistance in UTIs, then all episodes fulfilling a case definition would be registered, and those demonstrated to have an ARO identified. The ARO rate would then be determined by dividing the number of cases associated with an ARO by the total number of cases. There are some important potential biases that may exist in this situation. Cases caused by AROs that are culture negative would lead to an underestimate of the true ARO rate; although in theory possible, demonstrating such a bias would be practically difficult to show because AROs are usually identified by culture. More importantly, however, is

the potential bias that may arise from surveillance that attempts to determine resistance rates in all UTIs by evaluating urine culture results alone [28]. This is because in all patients with UTIs, urine is more likely to be sent for culture in patients with complicated urinary tract anatomy, and prior known resistance and treatment failures who are more likely to have resistant organisms [29]. Thus the rate of ARO in patients who are cultured is not equally distributed, and culture-documented cases of UTI are likely to overestimate the true rate of antimicrobial resistance in all UTIs.

Another consideration with case definitions is that definitions may vary in surveillance studies/programmes and may lead to false interpretation of rates of resistance. For example, Folden *et al.* conducted a study by examining the charts of residents in a medical centre who had been diagnosed with methicillin-resistant *Staphylococcus aureus* (MRSA) [30]. The guidelines used for original diagnosis were the CDC guidelines for healthcare-associated MRSA [30]. The researchers reassessed the data using an alternative guideline, the healthcare-associated risk factors guidelines for MRSA, and found that community-acquired MRSA prevalence rates were 5% using the healthcare-associated risk factors criteria and 49% using the CDC guidelines over the same time period [30]. If the same criteria for defining organisms are not used in the various surveillance programmes there is a risk that the resulting resistance data may be not comparable.

### Case ascertainment

This bias is related to either incomplete identification of all episodes that meet a case definition within the surveillance population or a failure to exclude those that do not meet that definition. Means by which this case ascertainment bias may arise include selective surveillance, active *vs.* passive reporting, differential rates of test ordering, and inclusion of cases that are external to the population at risk. Bias may arise when surveillance is conducted in selected subsets of the population at risk. Traditionally, hospitals have been implicated as the primary source of resistant organisms. While this holds true in many cases, it should also be recognized that most studies of resistant organisms have been conducted in hospitalized patients such that they are inherently biased towards this conclusion. Extended-spectrum  $\beta$ -lactamase-producing (ESBL) *Escherichia coli* were previously

thought to be a predominantly hospital-acquired pathogen based on such hospital-based studies. However, when surveillance for ESBL *E. coli* was conducted in both hospitalized and community-based patients, it became readily apparent that the majority were community-onset cases, many related to international travel [31].

Another important consideration in ascertainment bias is whether an active or passive approach is taken. Active surveillance investigations apply some form of systematic search or mechanism to identify cases whereas passive surveillance relies on voluntary or routine reporting systems. Passive surveillance is an important risk for bias as it notoriously leads to decreased estimates of disease [32, 33]. Vergison *et al.* compared a period of passive surveillance followed by active surveillance for invasive pneumococcal disease and found that the rate of identification was twice as high in an active surveillance period compared to a passive surveillance period [32]. However, they did not report whether or not differences in antimicrobial susceptibilities occurred as a result. In another example, Modesitt *et al.* found that reporting for AIDS in a passive surveillance system was 36% lower than the number of cases found with retrospective active surveillance [33].

Differential test ordering occurs when differential rates of testing occur due to application of differing criteria or clinical practices. This may arise from testing protocols that may be implemented or as a result of individual clinician's practices. The most striking example of this in the performance of surveillance for asymptotically colonized cases. For example, implementation of a routine screening policy for all admissions for MRSA will inevitably lead to a higher rate of MRSA isolation than a risk-based or *ad-hoc* approach. Shannon & French discovered this in a programme screening for MRSA; because the investigators do not screen for methicillin susceptibility, the results are biased towards resistance if resistance rates are compared against a centre that does not actively screen for MRSA [34]. The magnitude of this bias as assessed by Shannon & French was found to be an increase of 6–10% in the resistance frequencies and 10–15% in the number of isolates per year, comparing active surveillance against passive surveillance [34]. Therefore, it is important to define isolates as either clinical specimens or screening specimens and to analyse them separately in order to compare the resistance rates of hospitals that have screening programmes and hospitals that do not [34].

The decision to submit cultures varies greatly between physicians [35]. Smellie *et al.* examined culturing rates in different hospitals in the UK and found 3- to 19-fold differences in the amount of testing when comparing the 10% of centres with the highest testing rates and the 10% of centres with the lowest testing rates [35]. Other studies have suggested that this may be more of a problem with non-invasive infections because a high percentage, if not all episodes, of severe invasive disease will be sent for culture; this is highly variable for surveillance cultures where the resistance rates may be dependent solely on the decision to culture [36]. Finally, there is also an inherent sampling bias possible in laboratory-based surveillance systems where clinicians are more likely to submit specimens for culture in cases of treatment failure which may result in inflated rates of resistance.

### Sampling bias

Hennekens & Buring describe sampling as the ability to ‘draw an inference about the experience of an entire population based on an evaluation of only a sample’ and ‘based on that estimate, we make an inference about the frequency of [disease] in the whole population’ [22]. Sampling bias arises when the sample differs in some systematic way from the larger population. One way to minimize this bias is to conduct population-based studies where all residents in the population at risk fulfilling a case definition are included. However, in many cases these designs are not practical, and sampling is needed. If random sampling is performed in an unbiased fashion then the sample should reflect the true population. However, if a sample of the cases occurring in the base population is taken without true random sampling, then there is a risk for bias to occur as a result of a failure to draw independent random samples from the population in relation to time, space, and location.

In multi-centred surveillance, bias may arise if participating centres are not randomly chosen from within all possible centres representing the population of interest. Typically large, tertiary-care, urban, referral hospitals that have higher rates of severe disease and antimicrobial resistance rates are more likely to participate (‘volunteer bias’) in these systems [37]. This applies to multi-national studies where selected centres participate; the results from these centres may not be representative of the overall country involved if regional differences exist within countries [37, 38]. One study found significant regional differences in

resistance rates with *Campylobacter* spp. infections and this trend was especially evident with resistance to fluoroquinolones; resistance rates were 35.9% in urban areas compared to 27.1% in rural areas [39]. While results might be reflective of similar locales, they do provide a biased assessment of the overall population.

While it is important to recognize that significant differences may be present *between* participating centres, it should also be noted that this bias may arise *within* participating centres. We previously conducted a study that looked at all urine culture submitted to a regional laboratory and compared the species distribution and antimicrobial susceptibility profiles within selected cohorts and then compared them with the overall dataset [40]. We found that species distributions and resistance rates varied dramatically with sampling during different time periods in the time of day, date, season, and year, with facilities, and between communities [40].

Many investigators have observed seasonal variations in antimicrobial resistance rates, and sampling restricted to one season may provide a biased assessment of the overall occurrence in the population [31, 39, 41]. Guevara *et al.* studied the effect of seasonality on the rate of otitis media pathogens and antimicrobial resistance in Costa Rica and found that penicillin-resistant and penicillin-intermediate strains of *Streptococcus pneumoniae* were more prevalent in the rainy season (39%) than in the dry season (18%) [42]. Another study performed in The Netherlands found that *Campylobacter* spp. infections resistant to fluoroquinolones and macrolides were lower in the summer season than in the winter season [39]. Pitout *et al.* found that there was significant seasonal variation in ESBL *E. coli* in Calgary, Canada [31]. Data from the European Antimicrobial Resistance Surveillance System (EARSS) indicates that the prevalence of penicillin non-susceptible *S. pneumoniae* is higher in the summer months although in this case the difference was not found to be significant; however, they were unable to account for these seasonal differences in the rates of penicillin non-susceptible *S. pneumoniae* [43]. Sampling during a particular season has significant potential to lead to biased assessments of overall ARO occurrence.

### Multiple counting

When a case is counted more than once for the same episode then bias may result. In ARO surveillance this

largely reflects the failure to remove duplicate isolates from the same episode of disease. Numerous publications have investigated the importance of duplicate elimination in the assessment of species distribution/occurrence and resistance rates [34, 44–50]. The inclusion of duplicates can lead to reporting increased resistance rates as resistant organisms have a higher probability of being isolated multiple times [20, 46]. Shannon & French found that inclusion of duplicates led to a modest increase in reported resistance rates of MRSA, and a much higher rate for gentamicin resistance in *Klebsiella pneumoniae* isolates, and furthermore, these differences varied depending on the duplicate exclusion period examined [34]. The magnitude of this bias was supported by several other studies that found yearly resistance rates for MRSA and other organisms were significantly affected by the inclusion of duplicates [44–46]. Both Cebrian *et al.* and Rodríguez *et al.* found that the application of time criteria for exclusion of duplicates had a significant bias in determining prevalence, but like Magee found that this did not greatly affect the determination of resistance [48–50]. Many investigators did find that the outcome of including duplicates had quite pronounced effects in certain organisms and little or no effect in others [34, 44]. In addition, Li *et al.* found that the magnitude of duplicate removal from in-patients had a significantly greater effect than duplicate removal from outpatients [45].

Using routine urine culture data from the Calgary Health Region, we found that using a duplicate exclusion criterion of 1 year, rates of resistance were significantly lower and the species distribution different than if duplicates were included [40]. There may be a bias associated with inclusion of the same episodes of disease as multiple cultures are often sent to the laboratory. Furthermore, Shannon & French found that there was a 10% difference in resistance rates when a 5-day limit for duplicate removal was compared with a 365-day limit for duplicate removal and Rodríguez *et al.* found that when the time period for eliminating isolates increased, the percentage of susceptible isolates also increased [34, 48].

### Laboratory practice and procedures

Policies that direct laboratory testing protocols can introduce bias into surveillance systems. These include non-standardized testing, selective testing, rule-based reporting, and inadequate species-level identification [35, 51–53]. Laboratories often have

different policies regarding which antibiotics to test isolates against and may only test certain antibiotics on an isolate if it is resistant to the first-line antibiotics [20, 51]. Researchers found that when comparing different laboratories, the proportion of isolates tested varied from 20% to 90%, showing a marked difference in testing policies. In addition, laboratories often do not completely identify some isolates and therefore, data regarding resistance may be pooled, leading to biased data [20, 51, 53]. One study found that for urinary coliforms with pooled estimates of resistance, laboratories reported 7- to 13-fold increased rates of resistance compared to those without [51]. When the data are grouped together and resistant organisms are not completely identified, the restricted test menus can lead to biased resistance rates.

The accuracy of laboratory results used in surveillance systems has been questioned [52, 54]. A proficiency study testing the abilities of laboratories to identify emerging resistant organisms and reporting procedures conducted by the World Health Organization (WHO) found that only 20% of laboratories reported fully acceptable results, with most laboratories having trouble with only a few isolates [52]. This is consistent with other studies that have confirmed laboratories often have difficulties isolating only a few resistant organisms [54, 55]. For example, Heginbotham and colleagues reported on around 300 000 routine community isolates from 14 laboratories in Wales and found that selective testing policies and incomplete species-level testing were associated with falsely elevated levels of resistance [51]. In one example from this study, the rates of ciprofloxacin resistance in *Haemophilus influenzae* isolates were 0.2% with non-selective testing and 1.8% with selective testing [51].

Standardization can affect the validity of surveillance data. Standardization of laboratory policies and procedures can be important in reducing the associated biases with surveillance data [53]. Some researchers argue that an international standard for antimicrobial susceptibility testing would improve the accuracy and comparability of results [53]. One study found that there was significant variation in resistance rates for resistance to ciprofloxacin and concluded that this is due to there being laboratory–laboratory variation in the definition of resistance [19]. These researchers emphasized that for ciprofloxacin, small variations in definitions of resistance can have major effects of resistance rates [19]. In studies, a centralized laboratory with explicit techniques, quality control,

and interpretation of results is needed to minimize the risk for laboratory-related biases.

## DISCUSSION AND CONCLUSION

In this paper we identified and explored six main areas of potential bias related to ARO surveillance systems. It is important to note that there are inevitably other biases or potential for errors in interpretation of surveillance data that we did not identify and discuss. These include issues surrounding maintenance of data integrity and appropriate statistical analysis. If there are errors in the statistical analysis or study design, inferences based on the data may be inaccurate [22]. Other programmatic issues that are not related to bias *per se* but that are important aspects of surveillance systems include reporting timeliness, cost, responsiveness, and feasibility. It is important to recognize that in many cases that we used in this report to illustrate potential biases are issues related to interpretation and generalization of results rather than necessarily methodological flaws. Our emphasis in this report is on identifying that surveillance methodology is a complex issue and that surveillance data must be interpreted in light of the methodology utilized.

There are a number of surveillance systems in use at present that demonstrate careful methodological design and minimize potential biases. A widely regarded example of a 'gold standard' bacterial surveillance system is the CDC Division of Bacterial and Mycotic Diseases Active Bacterial Core Surveillance ABCs programme [56]. The ABCs programme determines the incidence, epidemiological characteristics, and microbiology of invasive disease due to a number of selected bacterial pathogens in several large populations in the USA (total population about 34.2 million). Surveillance is active and all laboratories in the populations under surveillance participate such that sampling bias is minimized. Population census data from the surveillance regions are used as the denominator data. A case definition for invasive disease is used that is based on the isolation of a pathogen from a normally sterile body site. Only cases in residents of the base population are included, only first isolates are included per episode of clinical disease, and samples are referred to a central laboratory for confirmation. While multiple geographical regions of the USA are represented, to our knowledge they were not randomly selected from the overall USA population.

It is our opinion that surveillance reports should routinely identify potential limitations and issues surrounding generalization. There has been a movement towards a more standardized reporting of observational medical research such as with the STROBE statement [57]. It is our contention that antimicrobial resistance surveillance reports would benefit from a standard means of reporting with areas of methodological limitations, interpretation, and generalization of results highlighted. The biases discussed in this review do have the potential to greatly affect the validity and interpretation of surveillance data. However, the extent to which these biases affect existing reports in the literature remains to be determined and is a topic for future research.

## DECLARATION OF INTEREST

None.

## REFERENCES

1. **Conly J.** Antimicrobial resistance in Canada. *Canadian Medical Association Journal* 2002; **167**: 885–891.
2. **Goossens H.** European status of resistance in nosocomial infections. *Chemotherapy* 2005; **51**: 177–181.
3. **Marchese A, Schito G.** Role of global surveillance in combating bacterial resistance. *Drugs* 2001; **61**: 167–173.
4. **Verhoef J, Fluit A.** Surveillance uncovers the smoking gun for resistance emergence. *Biochemical Pharmacology* 2006; **71**: 1036–1041.
5. **Wilton P, et al.** Strategies to contain the emergence of antimicrobial resistance: a systematic review of effectiveness and cost-effectiveness. *Journal of Health Services Research and Policy* 2002; **7**: 111–117.
6. **Simonsen G, et al.** The antimicrobial resistance containment and surveillance approach – a public health tool. *Bulletin of the World Health Organization* 2004; **82**: 928–934.
7. **Critchley I, Karlowsky J.** Optimal use of antibiotic resistance surveillance systems. *Clinical Microbiology and Infection* 2004; **10**: 502–511.
8. **Thacker S, Berkelman R.** Public health surveillance in the United States. *Epidemiology Review* 1988; **10**: 164–190.
9. **Mulvey M, et al.** Molecular characterization of ceftaxime-resistant *Escherichia coli* from Canadian hospitals. *Antimicrobial Agents and Chemotherapy* 2005; **49**: 358–365.
10. **Rhomberg P, et al.** Clonal occurrences of multidrug-resistant Gram-negative bacilli: report from the Meropenem Yearly Susceptibility Test Information Collection Surveillance Program in the United States.

- Diagnostic Microbiology and Infectious Disease* 2004; **54**: 249–257.
11. **Jones R, et al.** An overview of the Meropenem Yearly Susceptibility Test Information Collection (MYSTIC) program: 1997–2004. *Diagnostic Microbiology and Infectious Disease* 2004; **53**: 247–256.
  12. **Kyaw M, et al.** Effect of introduction of the pneumococcal conjugate vaccine on drug-resistant *Streptococcus pneumoniae*. *New England Journal of Medicine* 2006; **354**: 1455–1463.
  13. **Styers D, et al.** Laboratory-based surveillance of current antimicrobial resistance patterns and trends among *Staphylococcus aureus*: 2005 status in the United States. In: *Annals of Clinical Microbiology and Antimicrobials* 2006; **5**: 2.
  14. **Bax R, et al.** Surveillance of antimicrobial resistance – what, how and whither? *Clinical Microbiology & Infection* 2001; **7**: 316–325.
  15. **Masterton R.** Surveillance studies: how can they help the management of infection? *Journal of Antimicrobial Chemotherapy* 2000; **46** (Suppl. B): 53–58.
  16. **Felmingham D.** The need for antimicrobial resistance surveillance. *Journal of Antimicrobial Chemotherapy* 2002; **50** (Suppl. S1): 1–7.
  17. **Tapsall J.** Monitoring antimicrobial resistance for public health action. *Communicable Disease Intelligence* 2003; **27** (Suppl.): S70–74.
  18. **Hunter P, Reeves D.** The current status of surveillance of resistance to antimicrobial agents: report on a meeting. *Journal of Antimicrobial Chemotherapy* 2002; **49**: 17–23.
  19. **Livermore DM, et al.** Are routine sensitivity test data suitable for the surveillance of resistance? Resistance rates amongst *Escherichia coli* from blood and CSF from 1991–1997, as assessed by routine and centralized testing. *Journal of Antimicrobial Chemotherapy* 2000; **45**: 205–211.
  20. **Cornaglia G, et al.** European recommendations for antimicrobial resistance surveillance. *Clinical Microbiology & Infection* 2004; **10**: 349–383.
  21. **Monnet D.** Toward multinational antimicrobial resistance surveillance systems in Europe. *International Journal of Antimicrobial Agents* 2000; **15**: 91–101.
  22. **Hennekens CH, Buring JE.** *Epidemiology in Medicine*. Boston: Lippincott Williams & Wilkins, 1987.
  23. **Garner J, et al.** *CDC Definitions for Nosocomial Infections*. St. Louis: Mosby, 1996.
  24. **Laupland KB, et al.** Intensive-care-unit-acquired bloodstream infections in a regional critically ill population. *Journal of Hospital Infection* 2004; **58**: 137–145.
  25. **Laupland KB, et al.** Community-onset urinary tract infections: a population-based assessment. *Infection* 2007; **35**: 150–153.
  26. **Schwaber MJ, De-Medina T, Carmeli Y.** Epidemiological interpretation of antibiotic resistance studies – what are we missing? *Nature Reviews Microbiology* 2004; **2**: 979–983.
  27. **Mehnert-Kay SA.** Diagnosis and management of uncomplicated urinary tract infections. *American Family Physician* 2005; **72**: 451–456.
  28. **McIsaac W, et al.** Community-acquired antibiotic resistance in urinary isolates from adult women in Canada. *Canadian Journal of Infectious Diseases and Medical Microbiology* 2006; **17**: 337–340.
  29. **Wilson M, Gaido L.** Laboratory diagnosis of urinary tract infections in adult patients. *Clinical Infectious Diseases* 2004; **38**: 1150.
  30. **Folden DV, et al.** Estimating the proportion of community-associated methicillin-resistant *Staphylococcus aureus*: two definitions used in the USA yield dramatically different estimates. *Journal of Hospital Infection* 2005; **60**: 329–332.
  31. **Pitout JD, et al.** Population-based laboratory surveillance for *Escherichia coli*-producing extended spectrum B-lactamases: importance of community isolates with blaCTX-M genes. *Clinical Infectious Diseases* 2004; **38**: 1736–1741.
  32. **Vergison A, et al.** Epidemiological features of invasive pneumococcal disease in Belgian children: passive surveillance is not enough. *Pediatrics* 2006; **118**: e801–e809.
  33. **Modesitt SK, Hulman S, Fleming D.** Evaluation of active versus passive AIDS surveillance in Oregon. *American Journal of Public Health* 1990; **80**: 463–464.
  34. **Shannon K, French G.** Antibiotic resistance: effect of different criteria for classifying isolates as duplicates on apparent resistance frequencies. *Journal of Antimicrobial Chemotherapy* 2002; **49**: 201–204.
  35. **Smellie W, Clark G, McNulty C.** Inequalities of primary care microbiology testing between hospital catchment areas. *Journal of Clinical Pathology* 2003; **56**: 933–936.
  36. **McNulty C, et al.** Laboratory diagnosis of urinary symptoms in primary care – a qualitative study. *Communicable Disease & Public Health* 2003; **6**: 44–50.
  37. **Laverdiere M, et al.** Susceptibility patterns of *Candida* species recovered from Canadian intensive care units. *Journal of Critical Care* 2007; **22**: 245–250.
  38. **Paterson DL, et al.** International prospective study of *Klebsiella pneumoniae* bacteremia: implications of extended-spectrum  $\beta$ -lactamase production in nosocomial infections. *Annals of Internal Medicine* 2004; **140**: 26–33.
  39. **van Hees B, et al.** Regional and seasonal differences in incidence and antibiotic resistance of *Campylobacter* from a nationwide surveillance study in The Netherlands: an overview of 2000–2004. *Clinical Microbiology & Infection* 2007; **13**: 305–310.
  40. **Laupland K, et al.** Investigation of sources of potential bias in laboratory surveillance for anti-microbial resistance. *Clinical & Investigative Medicine* 2007; **30**: E159–E166.
  41. **Michel P, et al.** Multi-Provincial *Salmonella* Typimurium Case-Control Study Steering Committee: Regional, seasonal, and antimicrobial resistance distributions of *Salmonella typimurium* in Canada: a multi-provincial study. *Canadian Journal of Public Health* 2006; **97**: 470–474.
  42. **Guevara S, et al.** Seasonal distribution of otitis media pathogens among Costa Rican children. *Pediatric Infectious Disease Journal* 2008; **27**: 12–16.



43. **EARSS Annual Report, 2001** ([http://www.rivm.nl/earss/Images/EARSS%20rapport%202001\\_tcm61-25026.pdf](http://www.rivm.nl/earss/Images/EARSS%20rapport%202001_tcm61-25026.pdf)). Accessed 14 January 2009.
44. **Lee S, et al.** Comparison of trends of resistance rates over 3 years calculated from results for all isolates and for the first isolate of a given species from a patient. *Journal of Clinical Microbiology* 2004; **42**: 4776–4779.
45. **Li F, et al.** Isolate removal methods and methicillin-resistance *Staphylococcus aureus* surveillance. *Emerging Infectious Diseases* 2005; **11**: 1552–1557.
46. **Horvat RT, et al.** Effect of duplicate isolates of methicillin-susceptible and methicillin-resistant *Staphylococcus aureus* on antibiogram data. *Journal of Clinical Microbiology* 2003; **41**: 4611–4616.
47. **Shannon K, French G.** Validation of the NCCLS proposal to use results only from the first isolate of a species per patient in the calculation of susceptibility frequencies. *Journal of Antimicrobial Chemotherapy* 2002; **50**: 965–969.
48. **Rodriguez JC, et al.** Criteria of time and antibiotic susceptibility in the elimination of duplicates when calculating resistance frequencies. *Journal of Antimicrobial Chemotherapy* 2003; **52**: 132–134.
49. **Magee J.** Effects of duplicate and screening isolates on surveillance of community and hospital antibiotic resistance. *Journal of Antimicrobial Chemotherapy* 2004; **54**: 155–162.
50. **Cebrian L, et al.** Influence of various criteria for elimination of duplicates when calculating the prevalence and antibiotic susceptibility of microorganisms associated with urinary infections. *International Journal of Antimicrobial Agents* 2005; **25**: 173–176.
51. **Heginbotham M, et al.** Laboratory testing policies and their effects on routine surveillance of community antimicrobial resistance. *Journal of Antimicrobial Chemotherapy* 2004; **53**: 1010–1017.
52. **Tenover F, et al.** Ability of laboratories to detect emerging antimicrobial resistance: proficiency testing and quality control results from the World Health Organization's external quality assurance system for antimicrobial susceptibility testing. *Journal of Clinical Microbiology* 2001; **39**: 241–250.
53. **White D, et al.** Antimicrobial resistance: standardisation and harmonisation of laboratory methodologies for the detection and quantification of antimicrobial resistance. *Revue Scientifique et Technique de l'Office International des Epizooties* 2001; **20**: 849–858.
54. **Hageman J, et al.** Antimicrobial proficiency testing of national nosocomial infections surveillance system hospital laboratories. *Infection Control & Hospital Epidemiology* 2003; **24**: 356–361.
55. **Bronzwaer S, et al.** Comparability of antimicrobial susceptibility test results from 22 European countries and Israel: an external quality assurance exercise of the European Antimicrobial Resistance Surveillance System (EARSS) in collaboration with the United Kingdom National External Quality Assurance Scheme (UK NEQAS). *Journal of Antimicrobial Chemotherapy* 2002; **50**: 775–777.
56. **CDC.** Active bacterial core surveillance (<http://www.cdc.gov/ncidod/dbmd/abcs/>). Accessed 22 November 2008.
57. **von Elm E, et al.** Strengthening the reporting of observational studies in epidemiology (STROBE) statement: guidelines for reporting observational studies. *British Medical Journal* 2007; **335**: 806–808.